

9. ANOVA

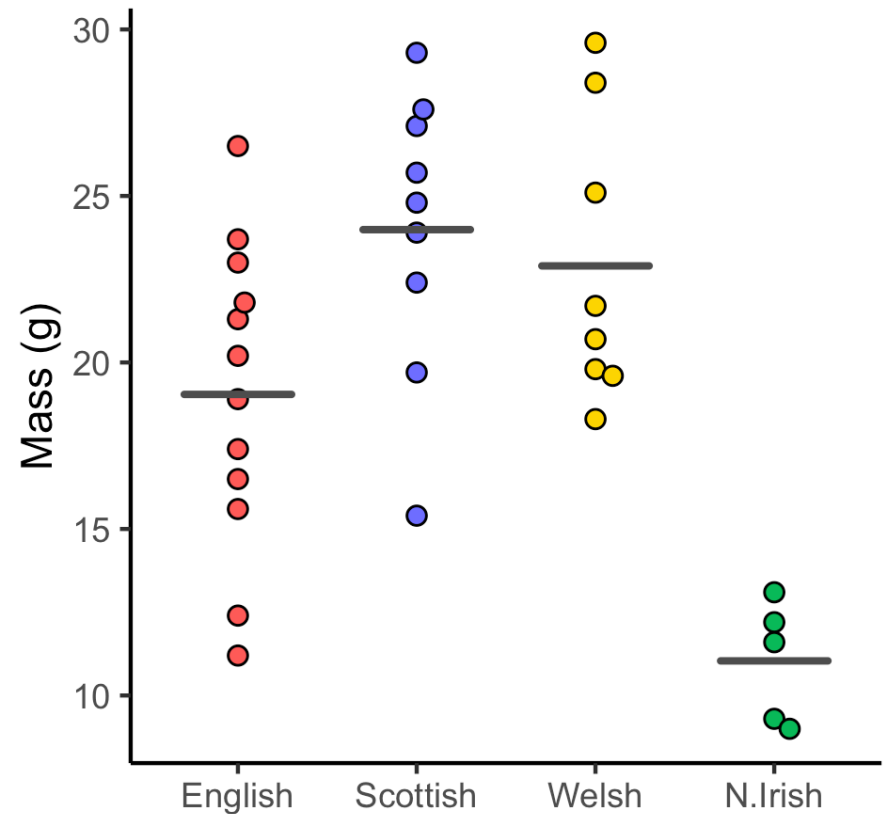
“Statistics is the science of variation”

Douglas M. Bates

One-way ANOVA

One-way ANOVA

- Extension of the t-test to more than 2 groups
- Null hypothesis: all samples came from populations with the same mean
- $H_0: \mu_1 = \mu_2 = \dots = \mu_k$
- The null hypothesis is tested by comparing variances
- ANOVA – **AN**alysis **Of** **V**ariance



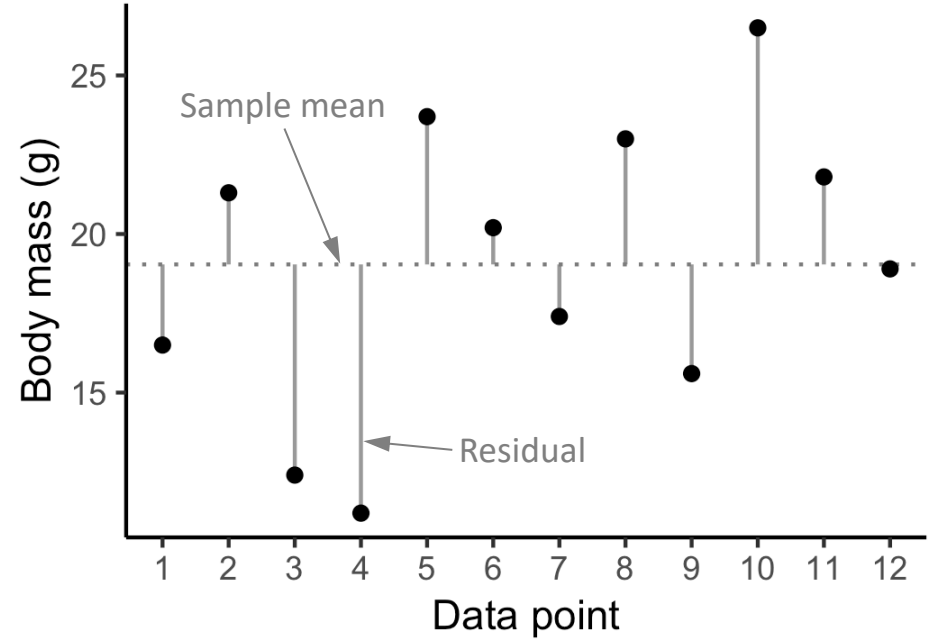
Variance

- Generalized variance: mean square

$$MS = \frac{SS}{\nu}$$

sum of squared residuals

number of degrees of freedom

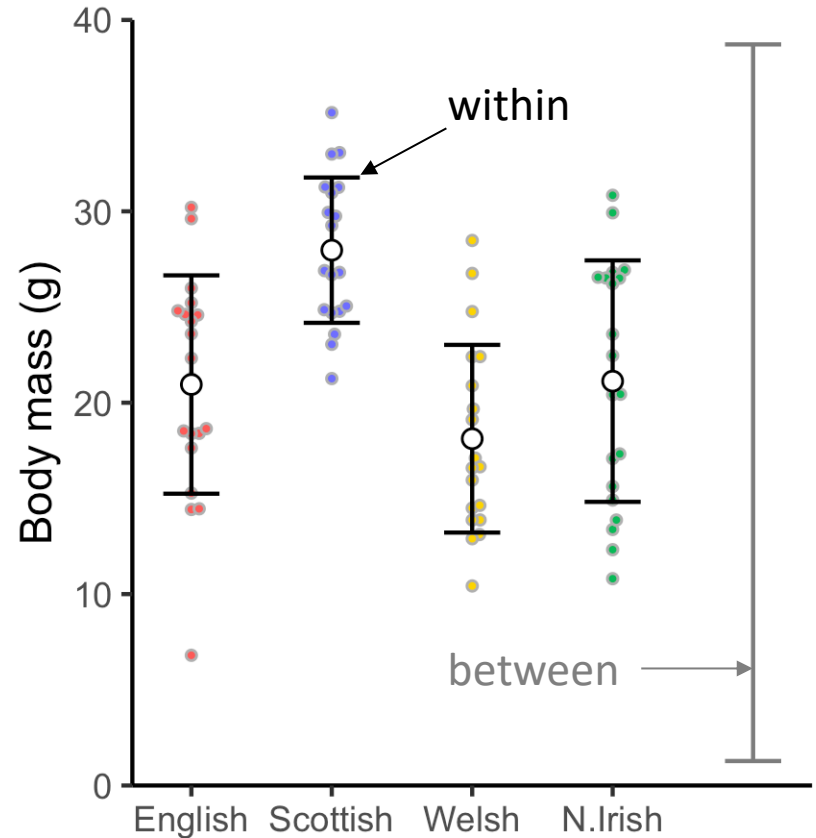
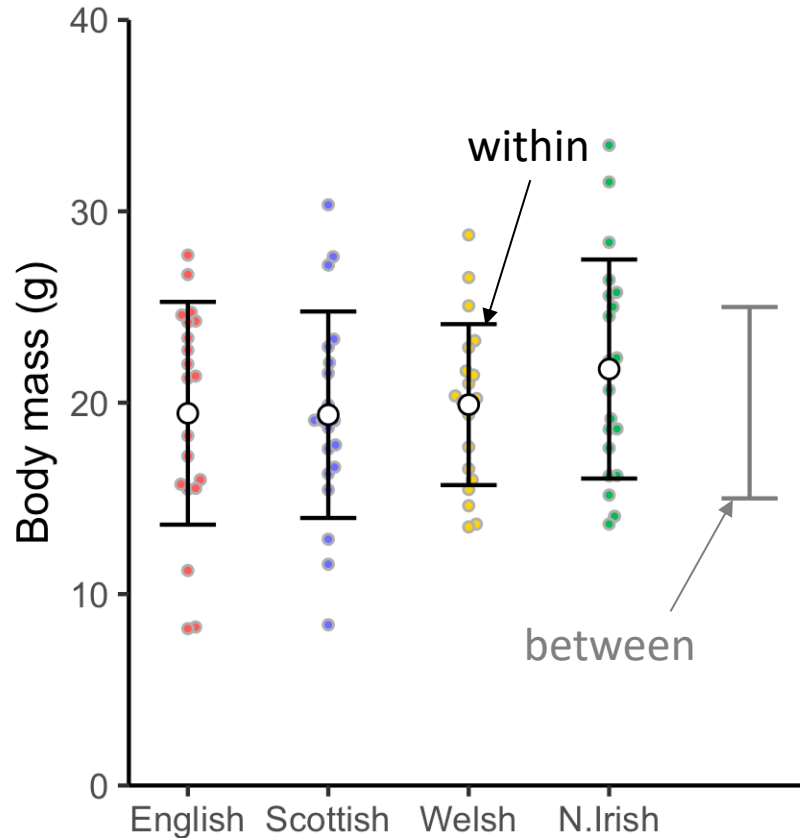


Variance represents spread of data around their mean

Variance = standard deviation²

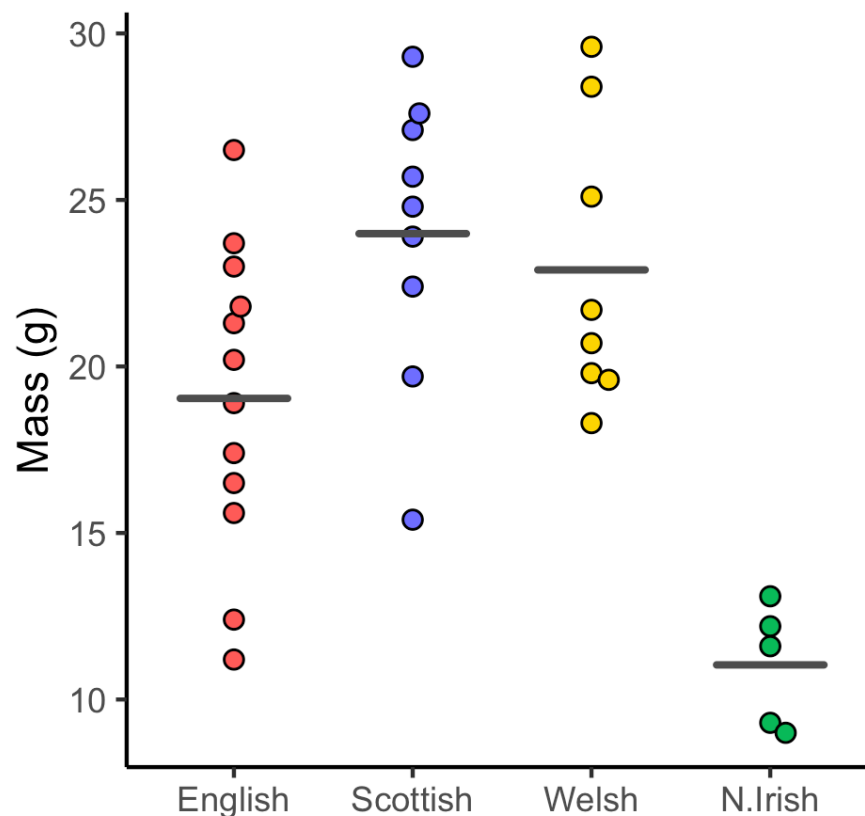
Variance between and within groups

- Variance within groups – typical variance in each group
- Variance between groups – how the sample mean varies from group to group



One-way ANOVA

- Null hypothesis: all samples came from populations with the same mean
- $H_0: \mu_1 = \mu_2 = \dots = \mu_k$
- **Assumption:** they all have common variance σ^2
- $n = 34$ data points
- $k = 4$ groups of data
- n_g - number of points in group g
- x_{gi} - body mass, group g , mouse i
- \bar{x}_g - mean in group g
- \bar{x} - grand mean, across all data points



Variance within groups

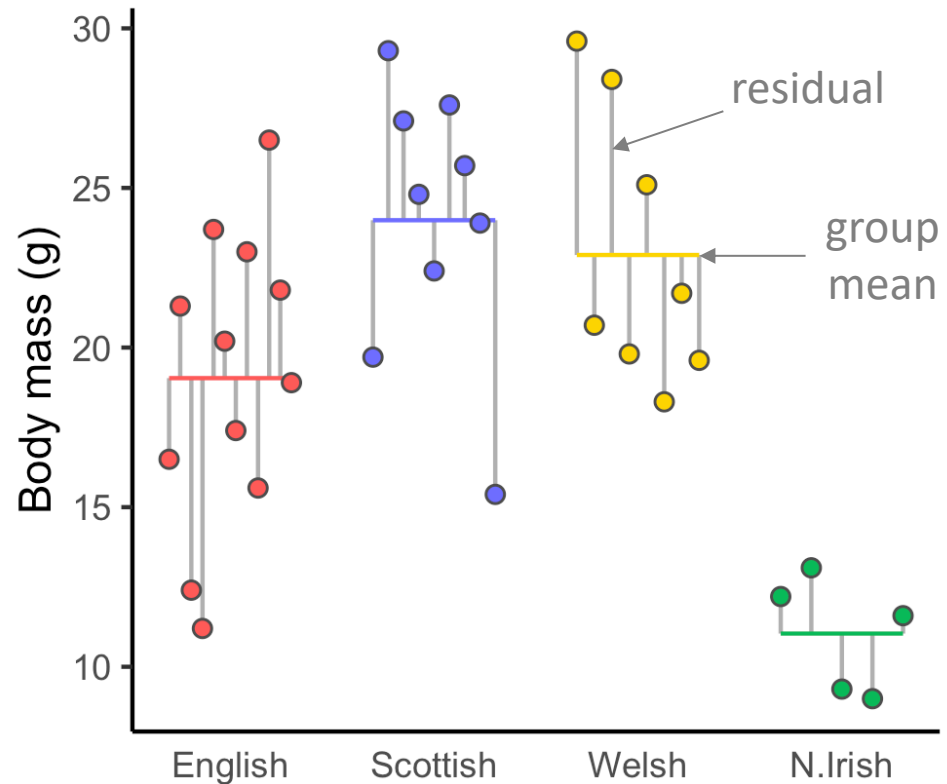
- Variance within groups is

$$MS_W = \frac{SS_W}{v_W}$$

sum of squared residuals

number of degrees of freedom

- MS_W estimates the common variance, σ^2 , regardless of the null hypothesis



mean of group g

$$SS_W = \sum_{g=1}^k \sum_{i=1}^{n_g} (x_{gi} - \bar{x}_g)^2$$

size of group g

$$v_W = \sum_{g=1}^k (n_g - 1)$$

Variance within groups

- Variance within groups is

$$MS_W = \frac{SS_W}{v_W}$$

Diagram illustrating the components of the within-group variance formula:

- SS_W is labeled "sum of squared residuals".
- v_W is labeled "number of degrees of freedom".

SS_W	524
v_W	30
MS_W	17.5

- MS_W estimates the common variance, σ^2 , regardless of the null hypothesis

mean of group g

$$SS_W = \sum_{g=1}^k \sum_{i=1}^{n_g} (x_{gi} - \bar{x}_g)^2$$

size of group g

$$v_W = \sum_{g=1}^k (n_g - 1)$$

Variance between groups

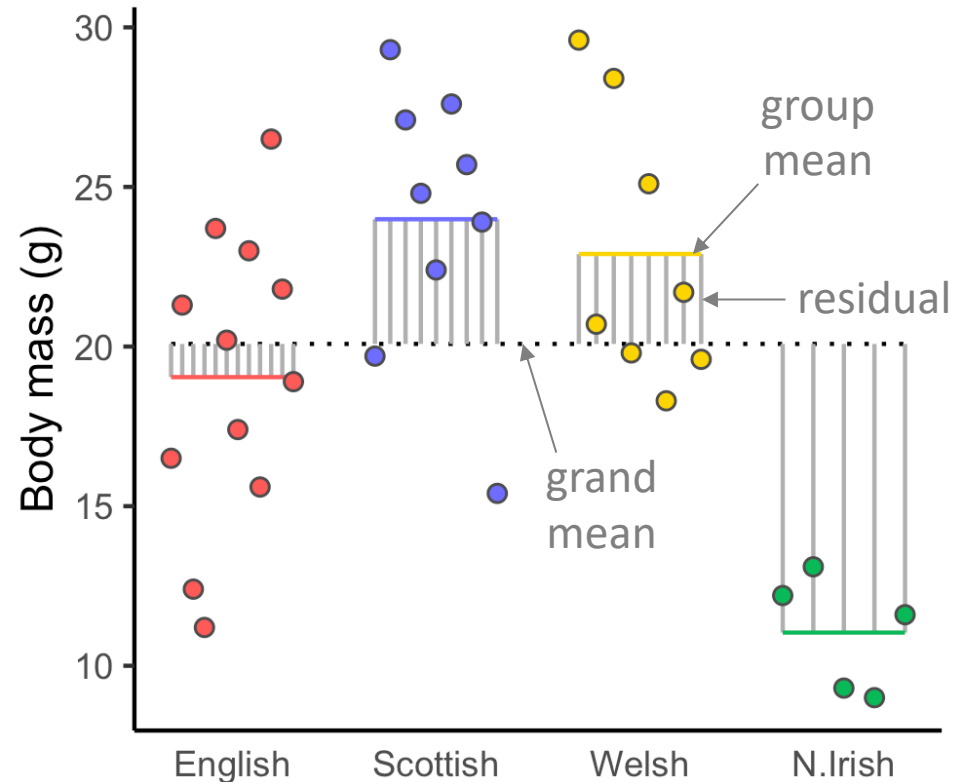
- Variance between groups is

$$MS_B = \frac{SS_B}{\nu_B}$$

sum of squared residuals

number of degrees of freedom

- MS_B estimates the common variance, σ^2 , **only** when the null hypothesis is true



grand mean

$$SS_B = \sum_{g=1}^k n_g (\bar{x}_g - \bar{x})^2$$

number of groups

$$\nu_B = k - 1$$

Variance between groups

- Variance between groups is

$$MS_B = \frac{SS_B}{\nu_B}$$

Diagram illustrating the components of the variance between groups formula:

- SS_B is labeled "sum of squared residuals".
- ν_B is labeled "number of degrees of freedom".

SS_W	524
ν_W	30
MS_W	17.5
SS_B	623
ν_B	3
MS_B	208

- MS_B estimates the common variance, σ^2 , **only** when the null hypothesis is true

grand mean

$$SS_B = \sum_{g=1}^k n_g (\bar{x}_g - \bar{x})^2$$

number of groups

$$\nu_B = k - 1$$

F-test

- MS_W estimates the common variance, σ^2 , regardless of the null hypothesis
- MS_B estimates the common variance, σ^2 , **only** when the null hypothesis is true

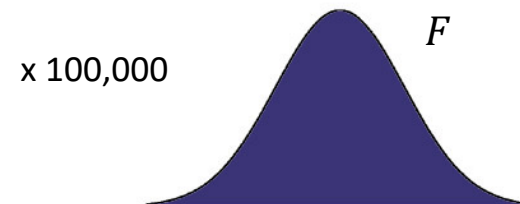
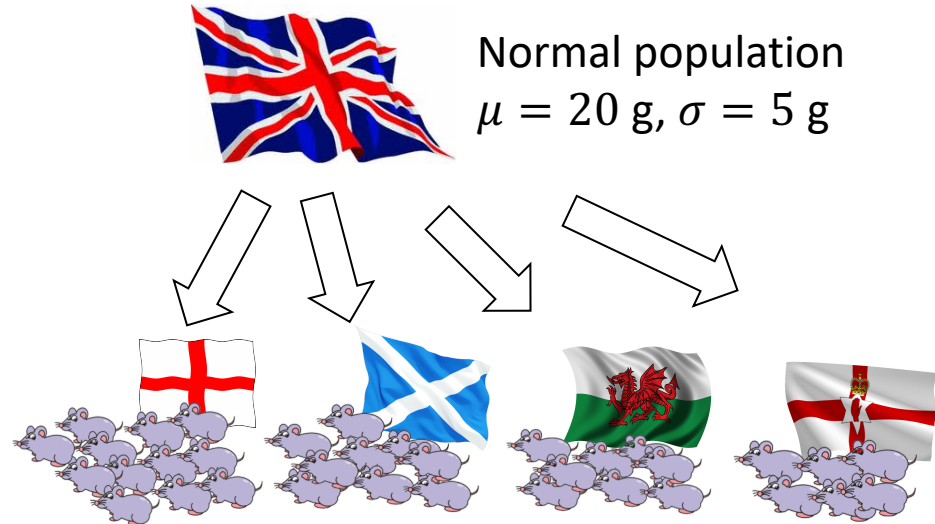
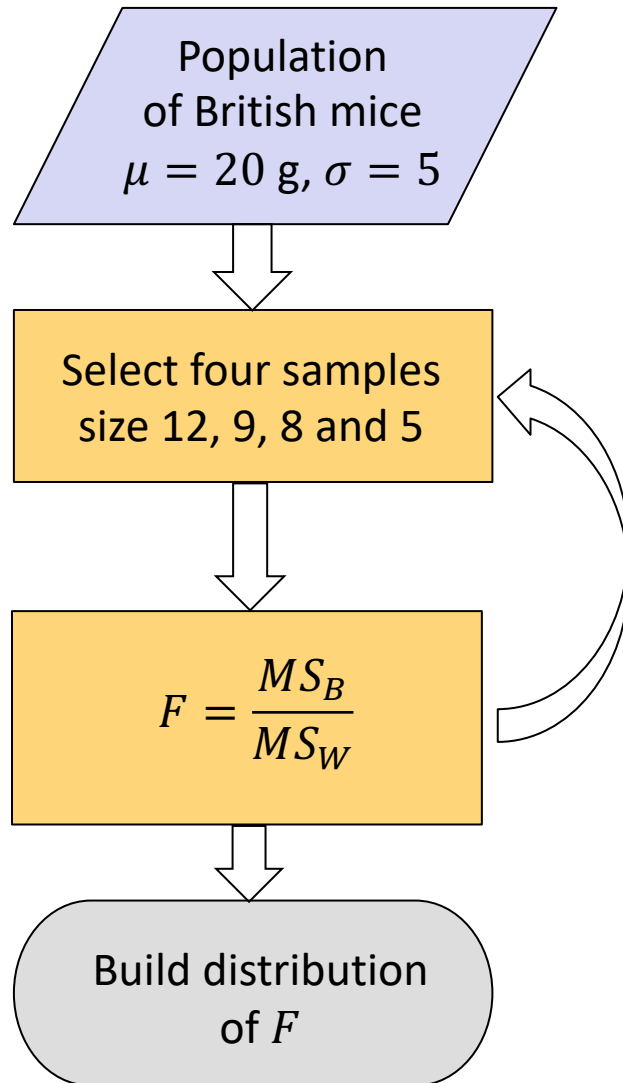
- Test for equality of variances: F-test

$$F = \frac{MS_B}{MS_W}$$

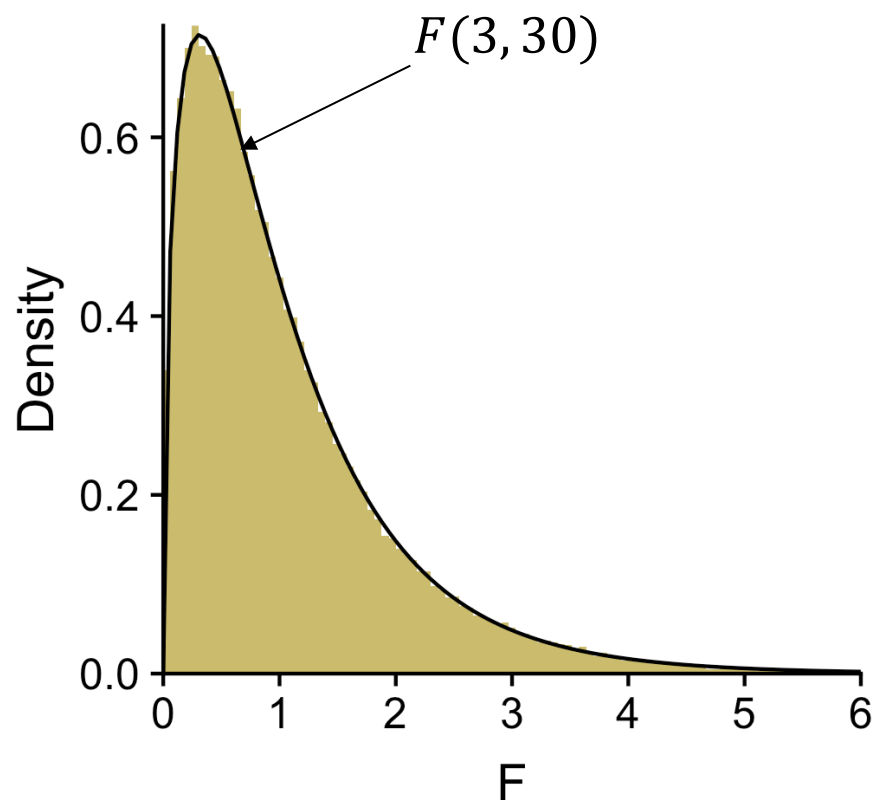
- Degrees of freedom: ν_B, ν_W
- If H_0 is true, we expect $F \sim 1$

SS_W	524
ν_W	30
MS_W	17.5
SS_B	623
ν_B	3
MS_B	208
F	11.9

Null distribution



Null distribution = F -distribution



Null distribution represents all random samples when the null hypothesis is true

SS_W	524
ν_W	30
MS_W	17.5
SS_B	623
ν_B	3
MS_B	208
F	11.9

```
> 1 - pf(11.9, 3, 30)  
[1] 2.648577e-05
```

ANOVA in R

```
# ANOVA
```

```
> mice <- read.table("http://tiny.cc/mice_1way", header=TRUE)
> mice.lm <- lm(Mass ~ Country, data=mice)
> anova(mice.lm)
```

Analysis of Variance Table

Response: Mass

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
Country	3	622.68	207.560	11.886	2.674e-05 ***
Residuals	30	523.89	17.463		

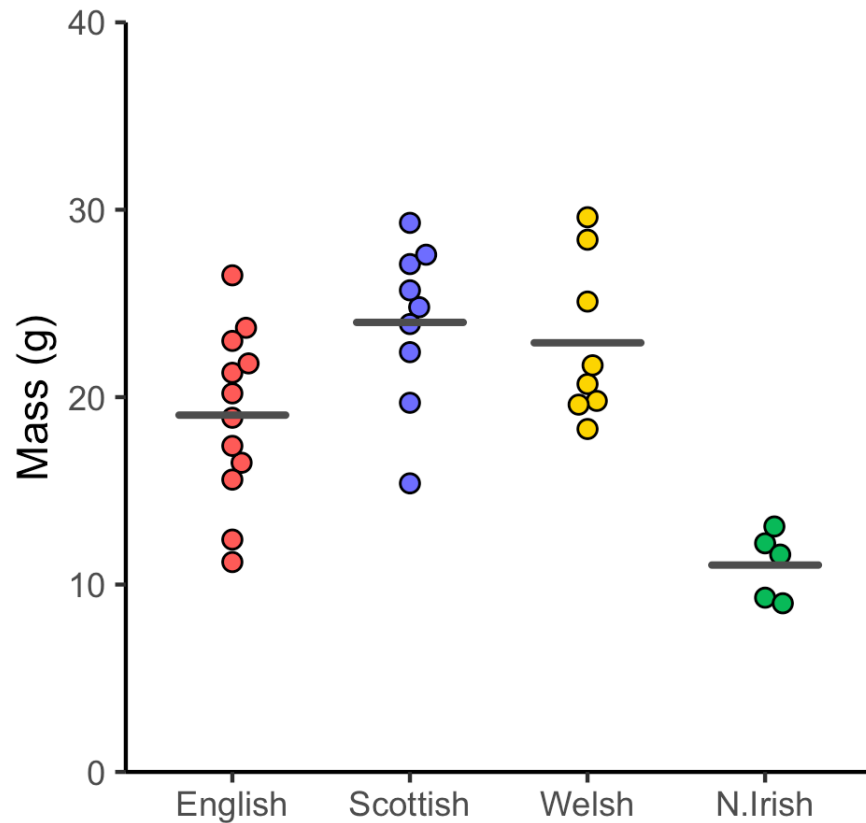
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```
> mice
```

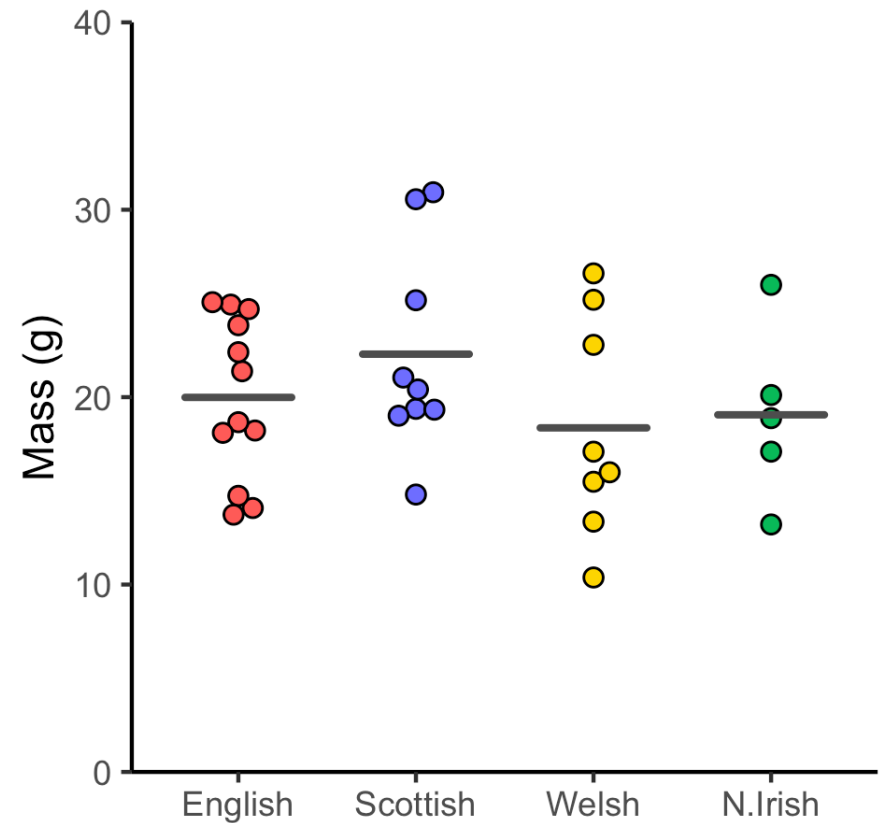
	Country	Mass
1	English	16.5
2	English	21.3
3	English	12.4
4	English	11.2
5	English	23.7
6	English	20.2
7	English	17.4
8	English	23.0
9	English	15.6
10	English	26.5
11	English	21.8
12	English	18.9
13	scottish	19.7
14	scottish	29.3
15	scottish	27.1
16	scottish	24.8
17	scottish	22.4
18	scottish	27.6
19	scottish	25.7
20	scottish	23.9
21	scottish	15.4
22	welsh	29.6
23	welsh	20.7
24	welsh	28.4
25	welsh	19.8
...		

Effect vs. no effect

MS_W 17.5 g²
 MS_B 208 g²
 F 11.9
 p 3×10^{-5}



MS_W 25.7 g²
 MS_B 24.3 g²
 F 0.94
 p 0.43



ANOVA assumptions

- Normality – data in each group are distributed normally
 - ANOVA is quite robust against non-normality
 - if strongly not normal (e.g. log-normal) – transform to normality
 - if this fails, use non-parametric Kruskal-Wallis test

- Independence – groups are independent
 - dependence: e.g., observations of the same subjects over time
 - if groups are not independent, ANOVA is not appropriate, use other methods

- Equality of variances – groups sampled from populations with the same variance
 - sometimes called homogeneity of variances, or homoscedasticity
/ˌhɒmɒʊskəˈdæstɪsɪti/
 - if variances are not equal, use Welch's approximated test

ANOVA and normality

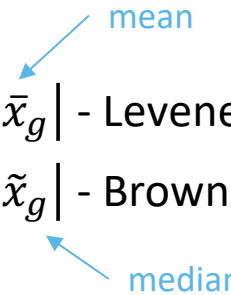
- My recommendation: do not perform normality test before ANOVA
- Transform data if possible

- Normality test (e.g. Shapiro-Wilk test)
 - underpowered for small samples (many false negatives)
 - oversensitive for large samples (many false positives)
 - nothing in nature is exactly normal!

- ANOVA is very robust to non-normality
 - Glass, G.V., P.D. Peckham, and J.R. Sanders. 1972. Consequences of failure to meet assumptions underlying fixed effects analyses of variance and covariance. *Rev. Educ. Res.* **42**: 237-288.
 - Harwell, M.R., E.N. Rubinstein, W.S. Hayes, and C.C. Olds. 1992. Summarizing Monte Carlo results in methodological research: the one- and two-factor fixed effects ANOVA cases. *J. Educ. Stat.* **17**: 315-339.
 - Lix, L.M., J.C. Keselman, and H.J. Keselman. 1996. Consequences of assumption violations revisited: A quantitative review of alternatives to the one-way analysis of variance F test. *Rev. Educ. Res.* **66**: 579-619.

Test to compare variances

- Null hypothesis: samples come from populations with equal variances
- $H_0: \sigma_1^2 = \sigma_2^2 = \dots = \sigma_k^2$
- Like ANOVA, except data x_{gi} are replaced by residuals R_{gi} :



The diagram consists of two blue arrows. The first arrow points from the word "mean" to the symbol \bar{x}_g in the equation for Levene's test. The second arrow points from the word "median" to the symbol \tilde{x}_g in the equation for Brown-Forsythe test.

$$R_{gi} = |x_{gi} - \bar{x}_g| \text{ - Levene's test}$$
$$R_{gi} = |x_{gi} - \tilde{x}_g| \text{ - Brown-Forsythe test}$$

- Test statistic:

$$W = \frac{MS_B}{MS_W}$$

Test to compare variances

- Null hypothesis: samples come from populations with equal variances
- $H_0: \sigma_1^2 = \sigma_2^2 = \dots = \sigma_k^2$
- Test statistic:

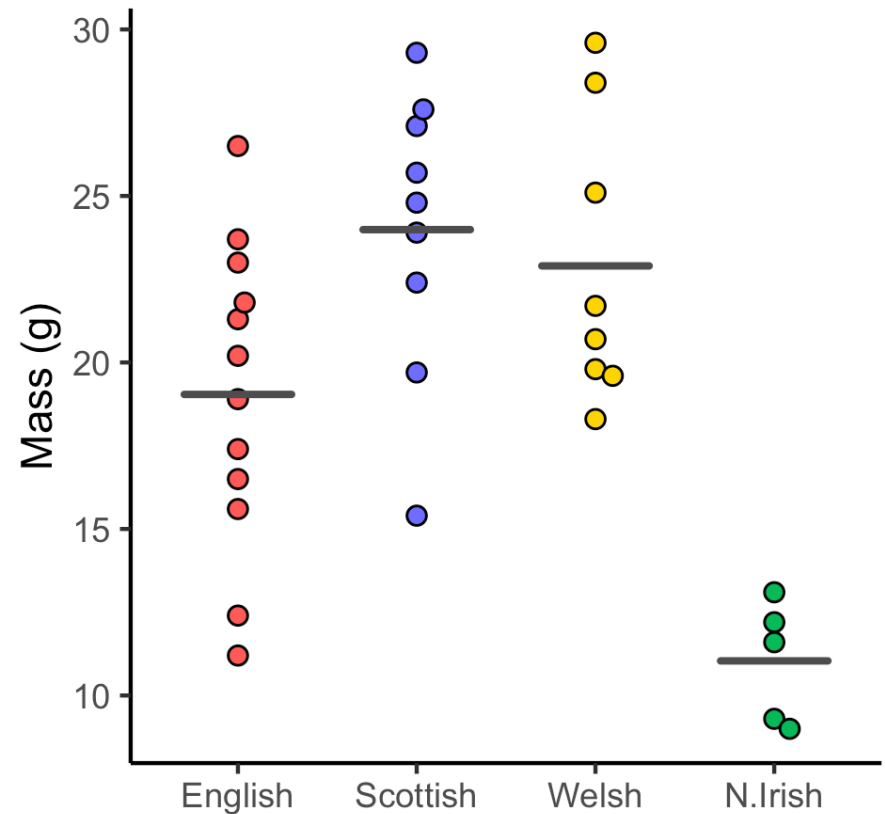
$$W = \frac{MS_B}{MS_W}$$

$$MS_B = 6.40 \text{ g}^2$$

$$MS_W = 6.89 \text{ g}^2$$

$$W = 0.930$$

$$p = 0.44$$



What if variances are not equal?

- B. L. Welch developed an approximated test
- Welch, B.L. (1951), “On the comparison of several mean values: an alternative approach”, *Biometrika*, **38**, 330–336
- Skip the details...

- Mice data

	F	ν_1	ν_2	p
ANOVA	11.89	3	30	2.7×10^{-5}
Welch's test	28.95	3	15.96	10^{-6}

Equality of variances in R

```
# Brown-Forsythe test for equality of variances
```

```
> library(lawstat)
```

```
> levene.test(mice$Mass, mice$Country)
```

modified robust Brown-Forsythe Levene-type test based on the absolute deviations from the median

```
data: mice$Mass
```

```
Test Statistic = 0.92948, p-value = 0.4386
```

```
# welch's test for unequal variances
```

```
> oneway.test(Mass ~ Country, mice, var.equal=FALSE)
```

One-way analysis of means (not assuming equal variances)

```
data: mass and country
```

```
F = 28.95, num df = 3.00, denom df = 15.96, p-value = 1.084e-06
```

Post-hoc analysis: Tukey's test

- A multiple t -test
- Finds differences and p-values for each pair of categories
- Post-hoc test, you need ANOVA first
- Skip the details...

	Scottish	Welsh	N.Irish
Welsh	-1.1 0.95		
N.Irish	-12.9 0.00003*	-11.9 0.0001*	
English	-4.9 0.05	-3.9 0.20	8.0 0.006*

Post-hoc test in R

```
# Tukey's Honest Significant Differences
```

```
> mice.av <- aov(mice.lm)
```

```
> TukeyHSD(mice.av)
```

```
Tukey multiple comparisons of means
```

```
95% family-wise confidence level
```

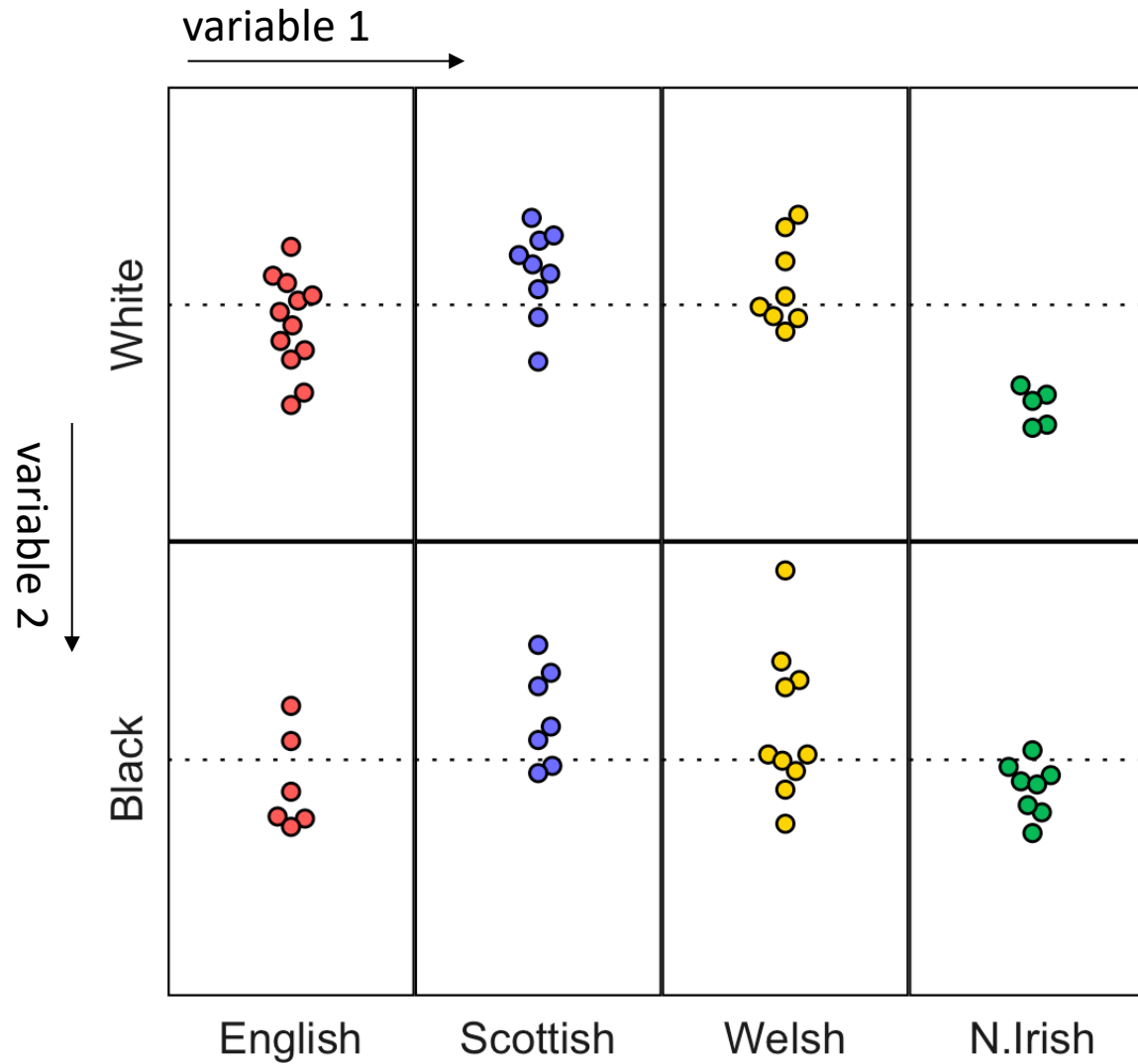
```
Fit: aov(formula = mice.lm)
```

```
$Country
```

	diff	lwr	upr	p adj
N.Irish-English	-8.001667	-14.04998948	-1.953344	0.0059422
Scottish-English	4.947222	-0.06331043	9.957755	0.0539580
Welsh-English	3.858333	-1.32806069	9.044727	0.2023039
Scottish-N.Irish	12.948889	6.61101070	19.286767	0.0000277
Welsh-N.Irish	11.860000	5.38219594	18.337804	0.0001394
Welsh-Scottish	-1.088889	-6.61022696	4.432449	0.9494897

Two-way ANOVA

Two variables



ANOVA as a linear model (one-way)

measurement group mean noise

$$x_{ir} = \mu_i + \epsilon_{ir}$$

group replicate

null hypothesis

$$H_0: \mu_1 = \mu_2 = \cdots = \mu_k$$

grand mean group effect

$$x_{ir} = \mu + \alpha_i + \epsilon_{ir}$$

null hypothesis

$$H_0: \alpha_1 = \alpha_2 = \cdots = \alpha_k = 0$$
$$\forall i: \alpha_i = 0$$

ANOVA as a linear model (two-way)

measurement

cell mean

noise

$$x_{ijr} = \mu_{ij} + \epsilon_{ijr}$$

column

row

replicate

grand mean

column effect

row effect

interaction effect

$$x_{ijr} = \mu + \alpha_i + \beta_j + \gamma_{ij} + \epsilon_{ijr}$$

Column means are equal:

$$H_0^{\text{col}}: \forall i: \alpha_i = 0$$

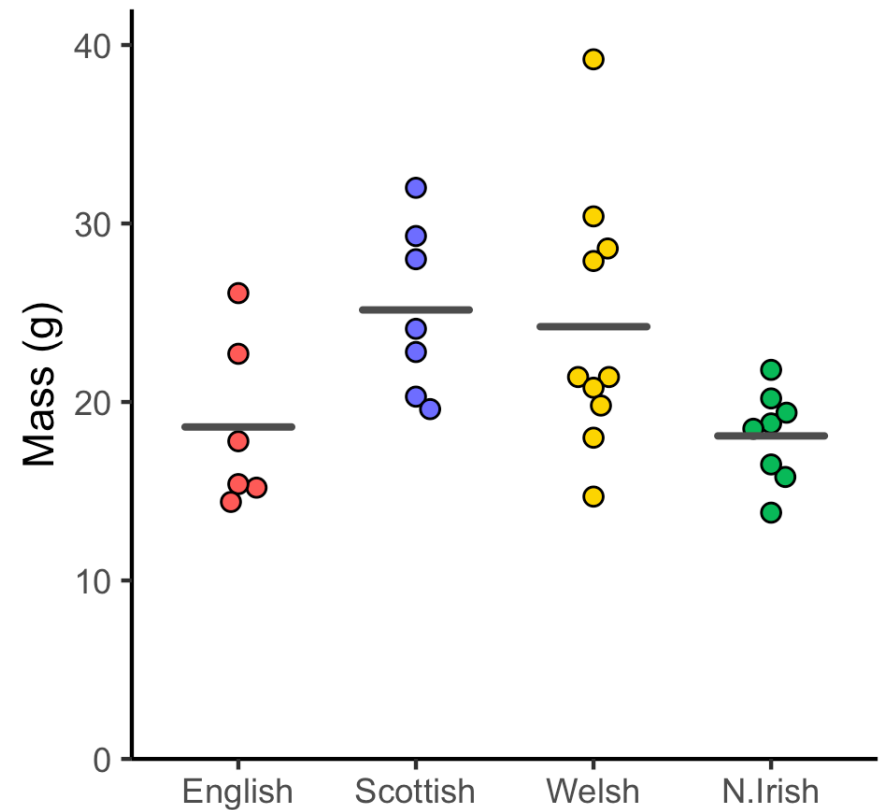
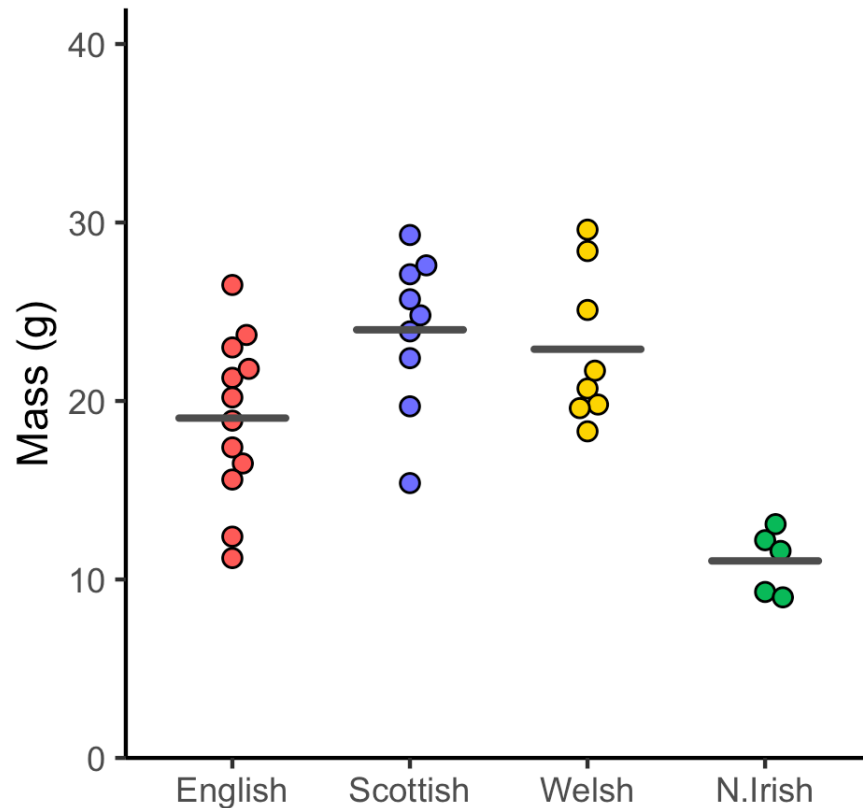
Row means are equal:

$$H_0^{\text{row}}: \forall j: \beta_j = 0$$

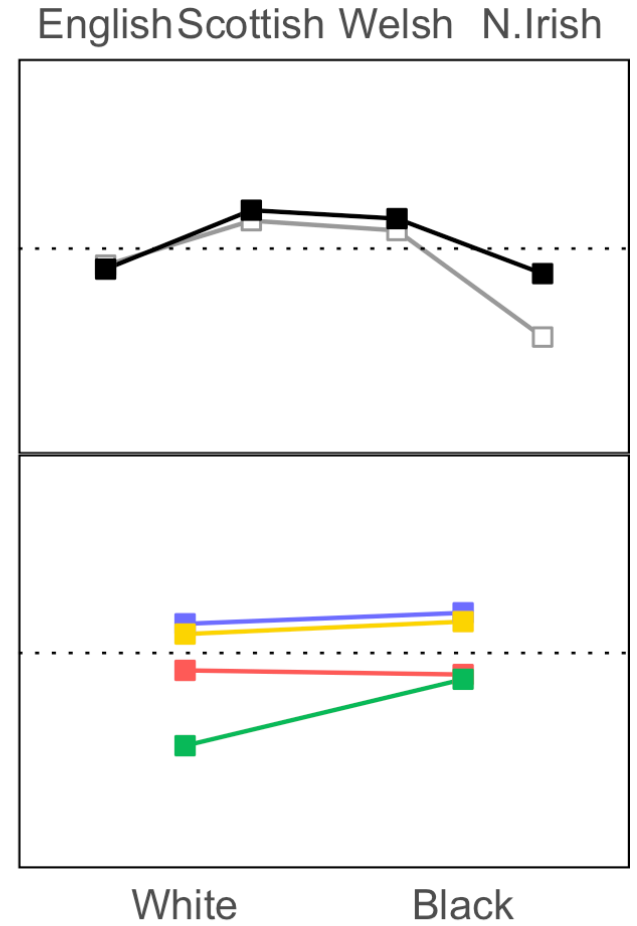
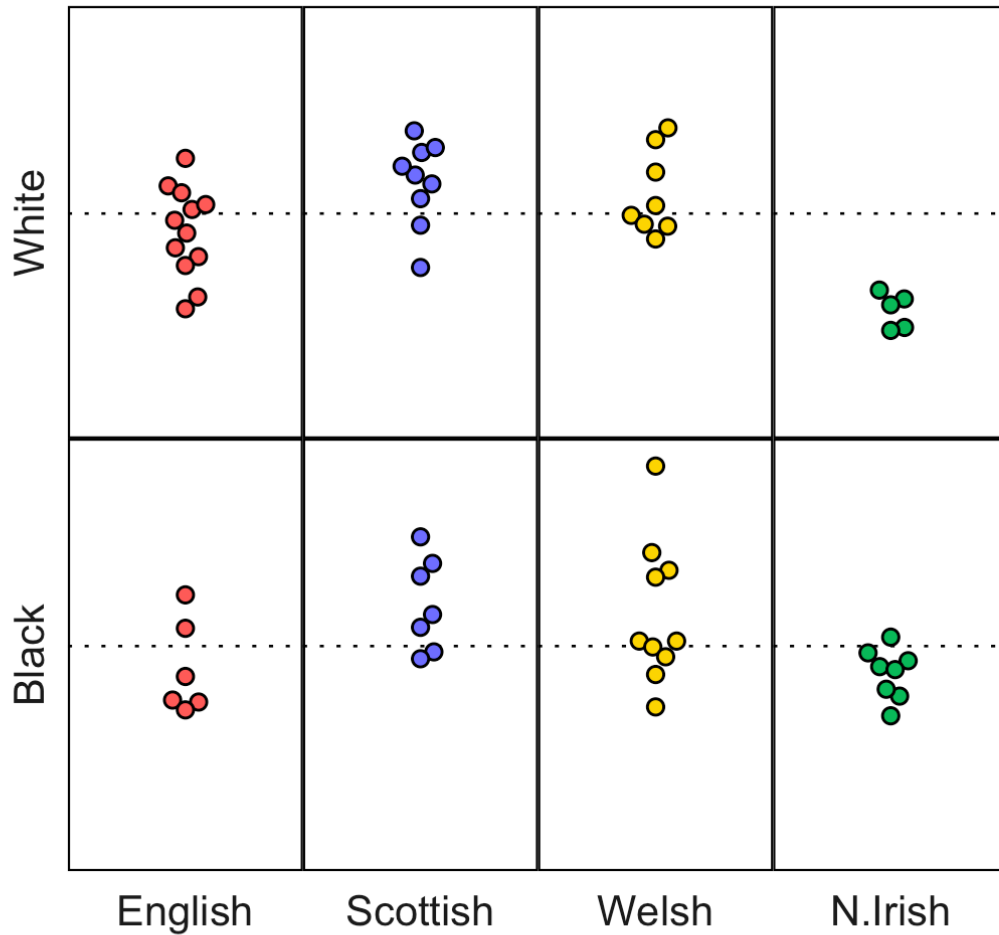
There is no interaction between rows and columns:

$$H_0^{\text{int}}: \forall i, j: \gamma_{ij} = 0$$

More mice!



Two-way ANOVA – two variables



How to do it in R?

```
# 2-way ANOVA
```

```
> mice <- read.table("http://tiny.cc/mice_2way", header=TRUE)
> mice.lm <- lm(Mass ~ Country + Colour + Country*Colour, data=mice)
> anova(mice.lm)
```

Analysis of Variance Table

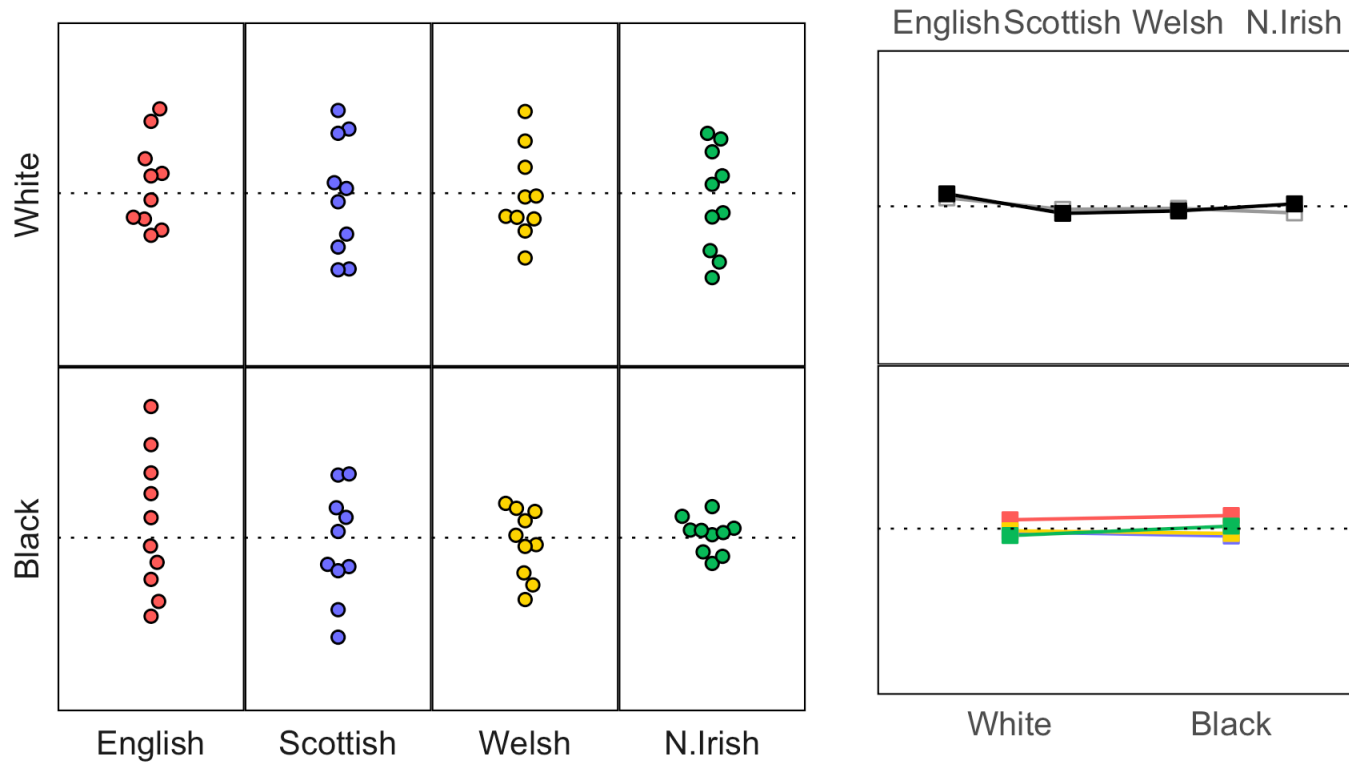
Response:

Mass	Df	Sum Sq	Mean Sq	F value	Pr(>F)
Country	3	809.68	269.893	11.9366	3.598e-06 ***
Colour	1	59.87	59.873	2.6480	0.1092
Country:Colour	3	107.39	35.797	1.5832	0.2034
Residuals	57	1288.80	22.611		

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1>

Null hypotheses: all three true

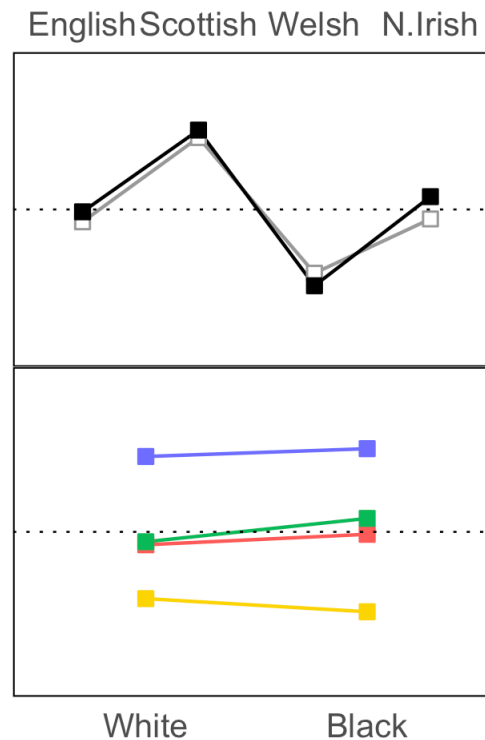
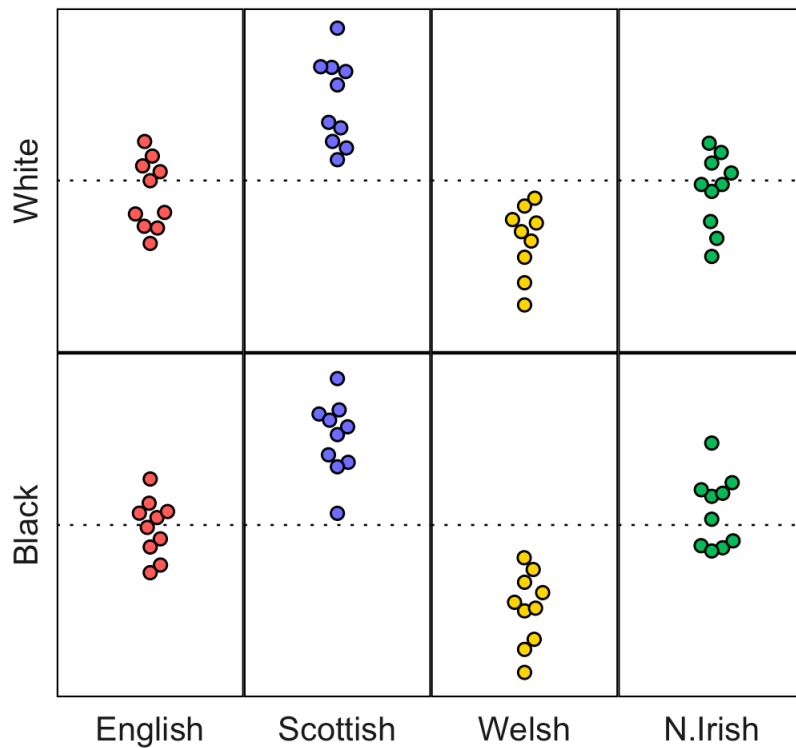
- $x_{ijr} = \mu + \alpha_i + \beta_j + \gamma_{ij} + \epsilon_{ijr}$
- $\mathbf{A} = \begin{pmatrix} 0 & 0 & 0 & 0 \end{pmatrix}, \mathbf{B} = \begin{pmatrix} 0 \\ 0 \end{pmatrix}, \mathbf{\Gamma} = \begin{pmatrix} 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 \end{pmatrix}$



	<i>p</i>
columns	0.68
rows	0.87
interaction	0.96

Null hypotheses: columns not equal

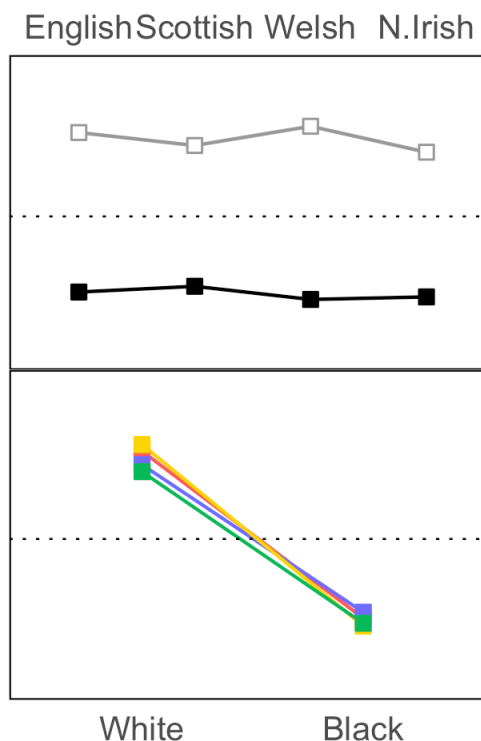
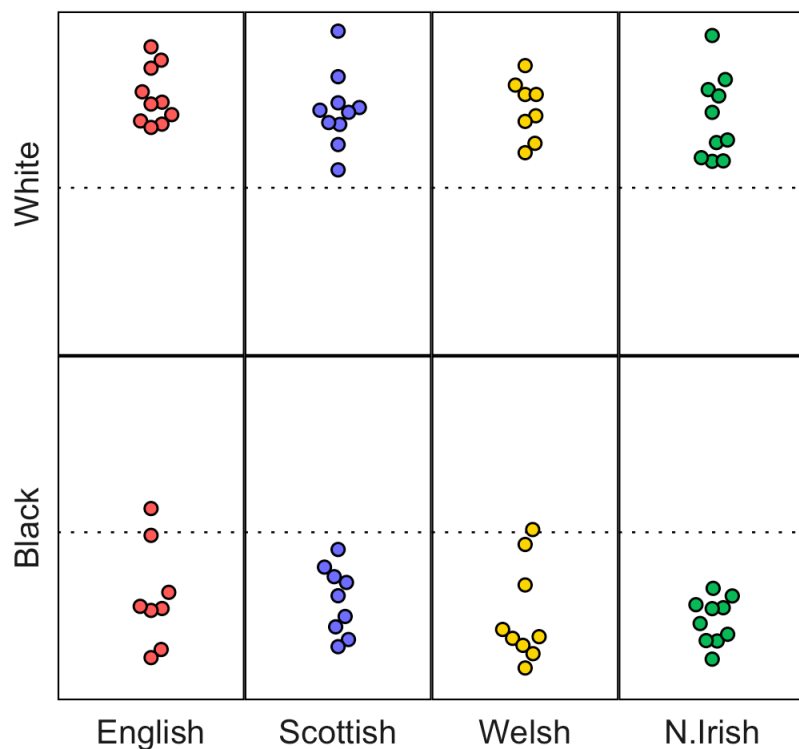
- $x_{ijr} = \mu + \alpha_i + \beta_j + \gamma_{ij} + \epsilon_{ijr}$
- $\mathbf{A} = \begin{pmatrix} 0 & 10 & -10 & 0 \end{pmatrix}, \mathbf{B} = \begin{pmatrix} 0 \\ 0 \end{pmatrix}, \mathbf{\Gamma} = \begin{pmatrix} 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 \end{pmatrix}$



	p
columns	2×10^{-19}
rows	0.38
interaction	0.46

Null hypotheses: rows not equal

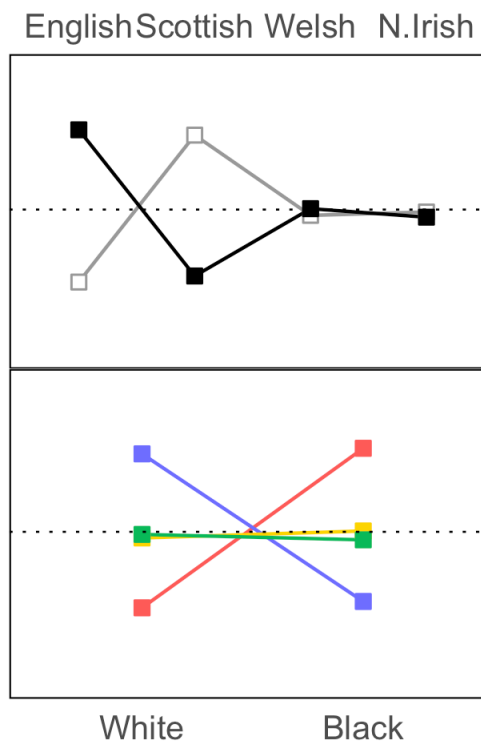
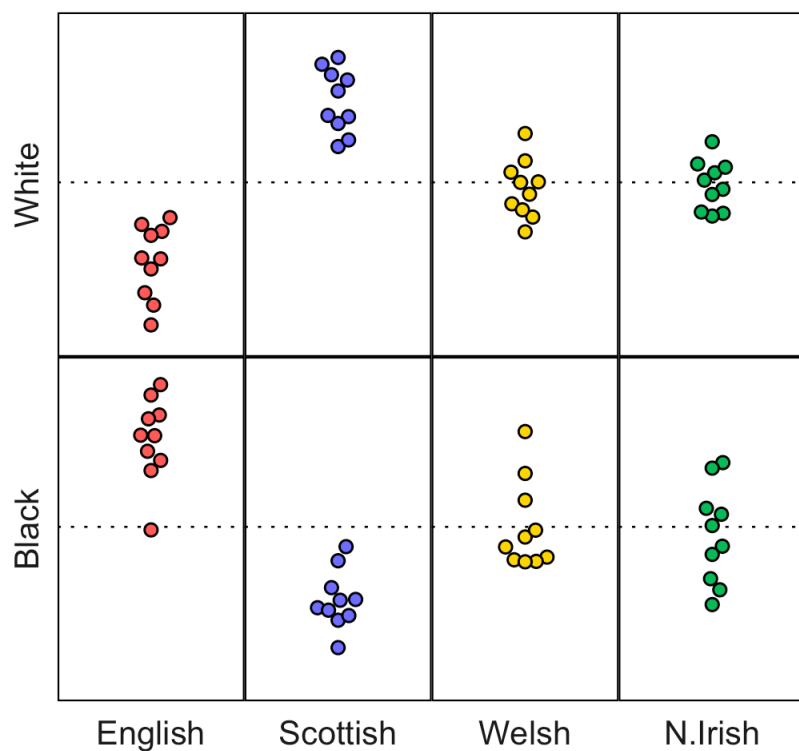
- $x_{ijr} = \mu + \alpha_i + \beta_j + \gamma_{ij} + \epsilon_{ijr}$
- $\mathbf{A} = \begin{pmatrix} 0 & 0 & 0 & 0 \end{pmatrix}$, $\mathbf{B} = \begin{pmatrix} 10 \\ -10 \end{pmatrix}$, $\mathbf{\Gamma} = \begin{pmatrix} 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 \end{pmatrix}$



	p
columns	0.75
rows	10^{-27}
interaction	0.56

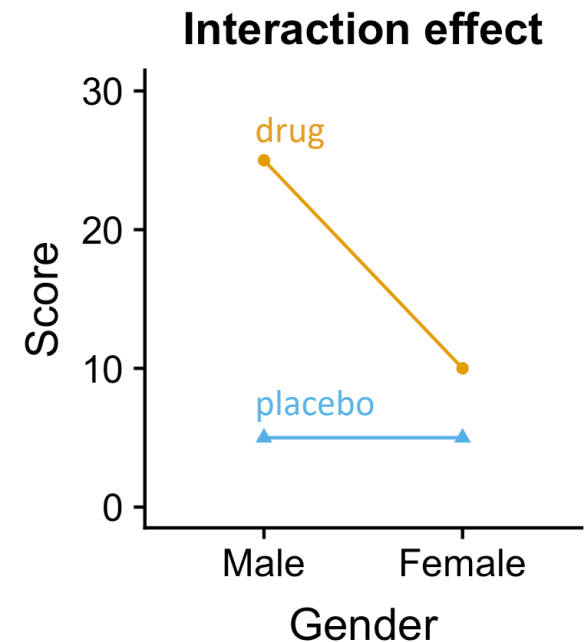
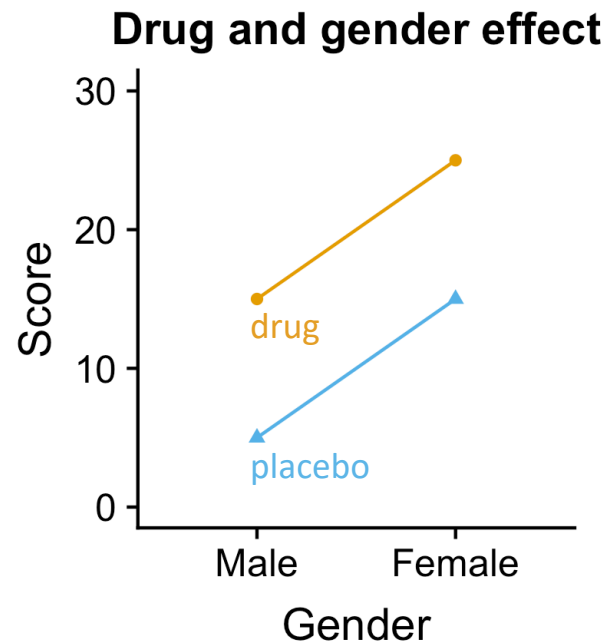
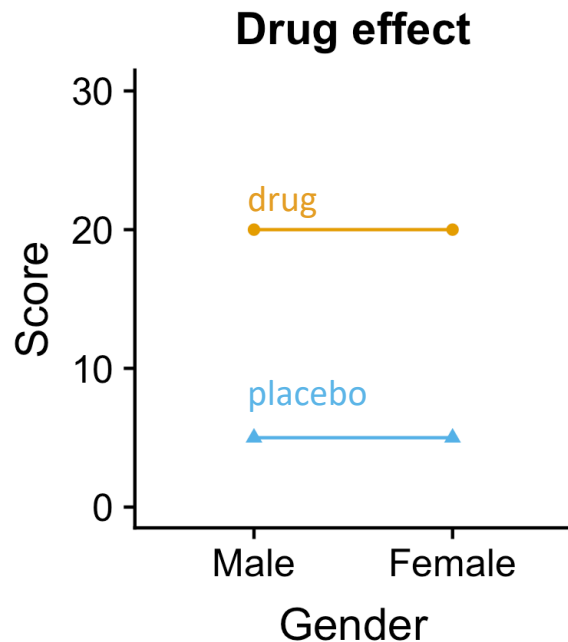
Null hypotheses: interaction

- $x_{ijr} = \mu + \alpha_i + \beta_j + \gamma_{ij} + \epsilon_{ijr}$
- $\mathbf{A} = \begin{pmatrix} 0 & 0 & 0 & 0 \end{pmatrix}, \mathbf{B} = \begin{pmatrix} 0 \\ 0 \end{pmatrix}, \boldsymbol{\Gamma} = \begin{pmatrix} -10 & 10 & 0 & 0 \\ 10 & -10 & 0 & 0 \end{pmatrix}$



	p
columns	0.78
rows	0.67
interaction	10^{-20}

Drug effect: three examples

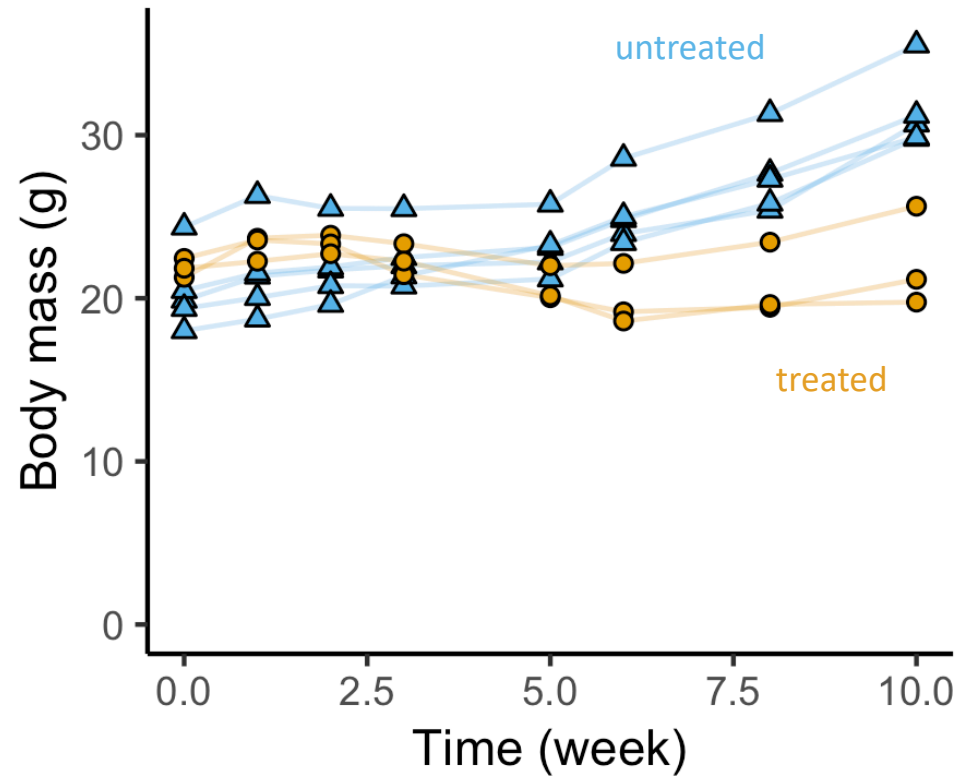


- male and female patients
- some were given drug, some were given placebo
- score measures response to the drug

Time-course experiments

Time-course experiments

- Obesity study in mice
- Two groups:
 - untreated
 - treated with a drug
- Feed them a lot
- Observe body mass over time
- Is there a difference between the two groups?



Time-course experiments

- You can do ANOVA
- $p = 5 \times 10^{-5}$

But

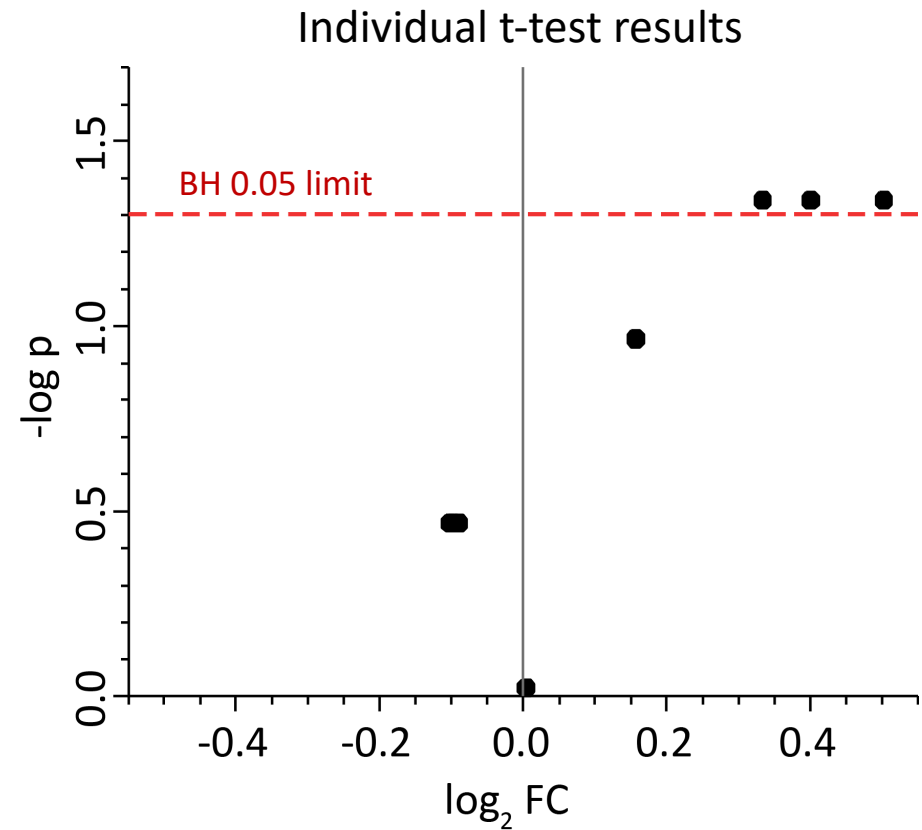
- Data are correlated
- ANOVA doesn't recognize numerical variables (time)

```
> dat <- read.table('http://tiny.cc/time_course', header=TRUE)
> dat.lm <- lm(Mass ~ Treatment + Time + Treatment*Time, dat)
> anova(dat.lm)
```

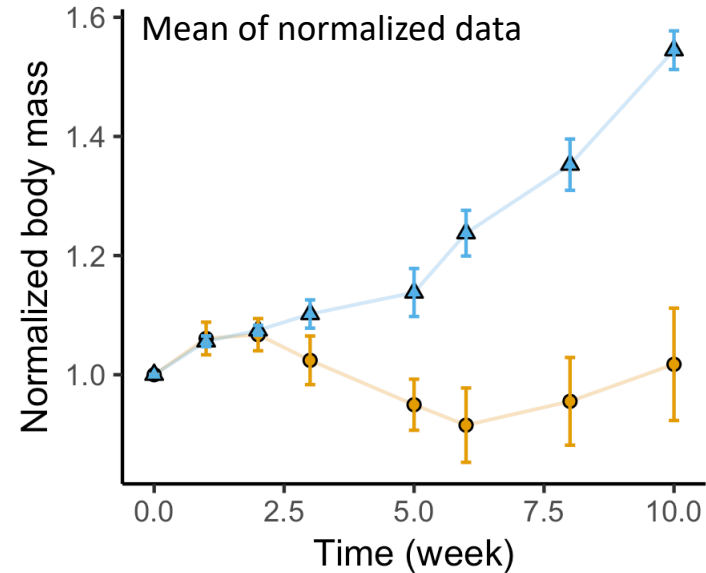
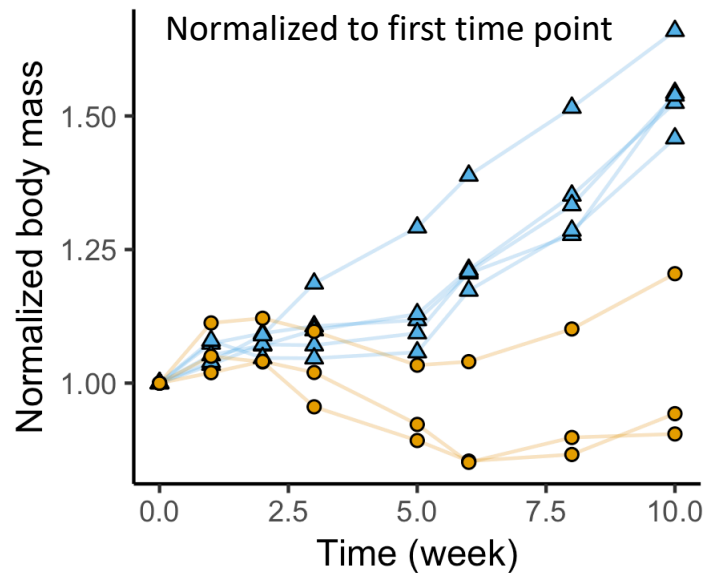
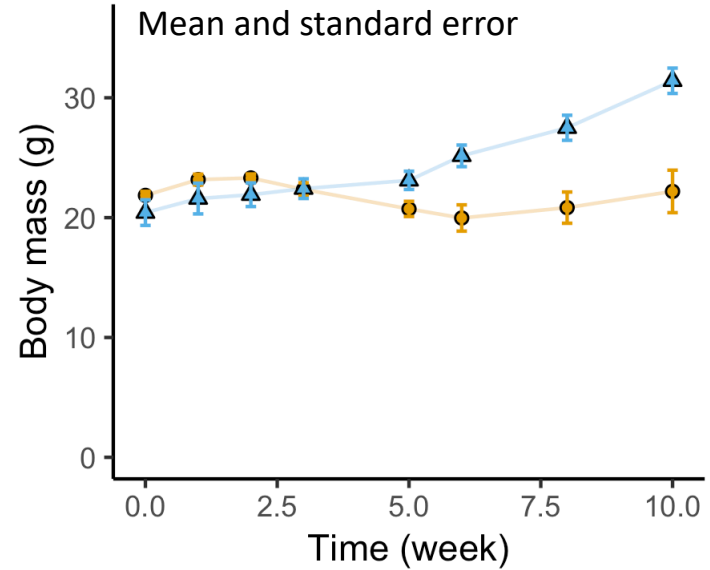
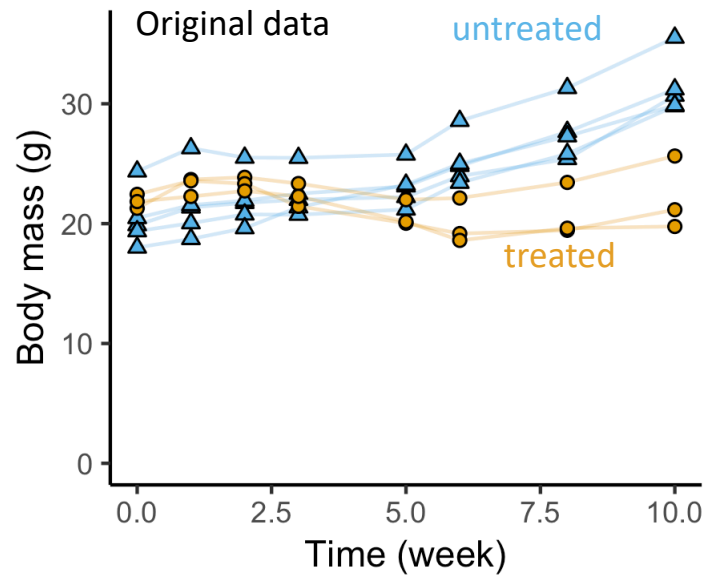
	Df	Sum Sq	Mean Sq	F value	Pr(>F)
Treatment	1	85.538	85.538	20.1508	4.481e-05
Time	7	272.465	38.924	9.1694	3.825e-07
Treatment:Time	7	230.738	32.963	7.7652	2.907e-06

Time-course experiments

- What about t-test at each time point?
- Works well!
- Three time points are significantly different
- But: misses point-to-point correlation

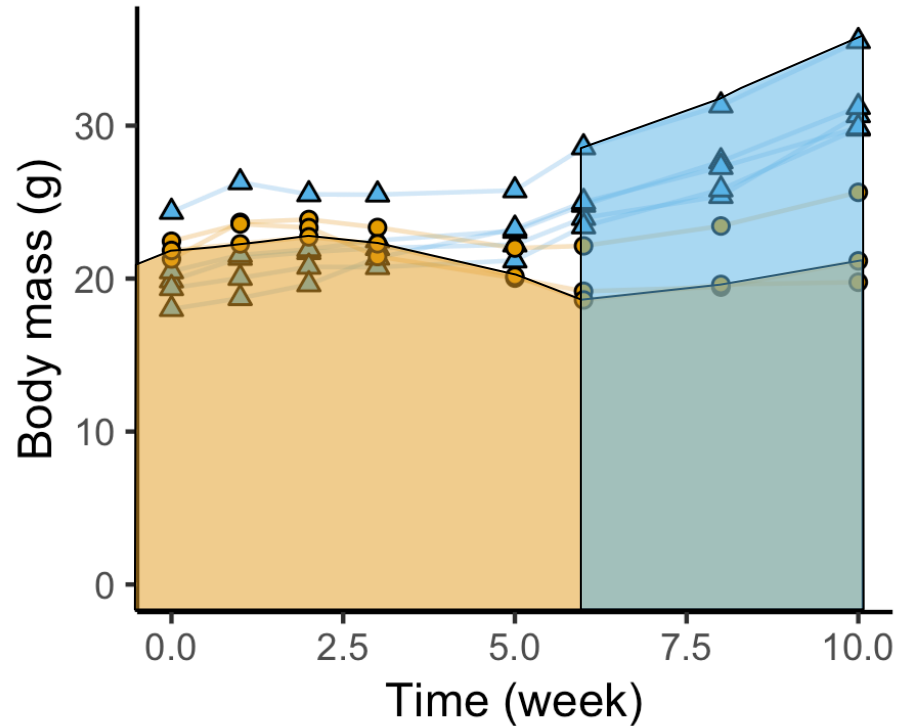


Data transformation

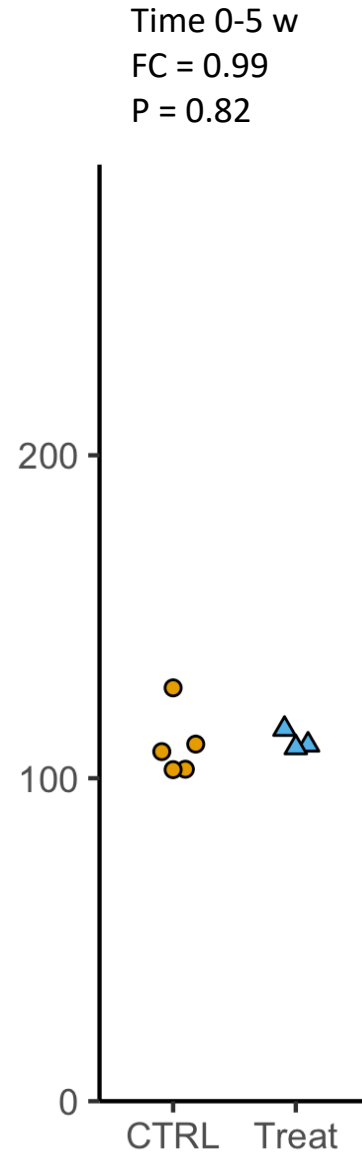
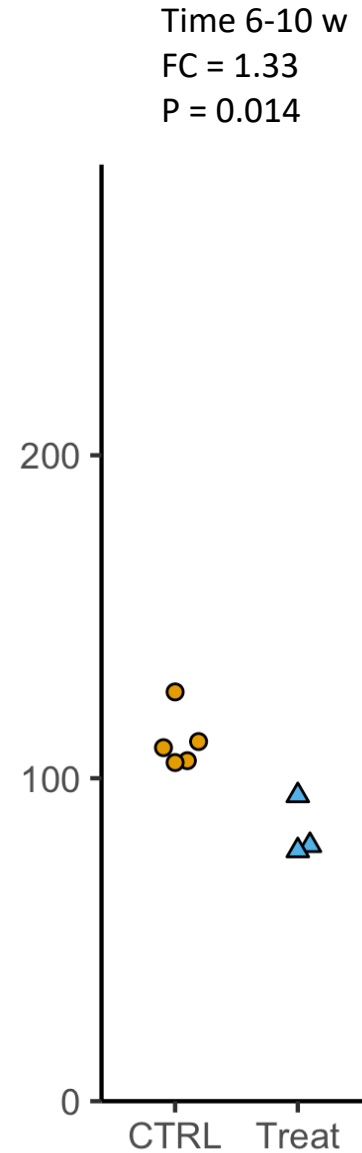
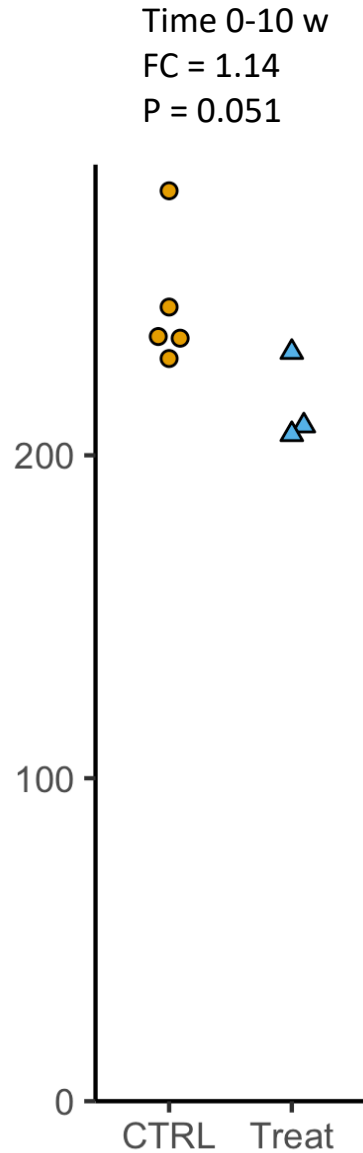
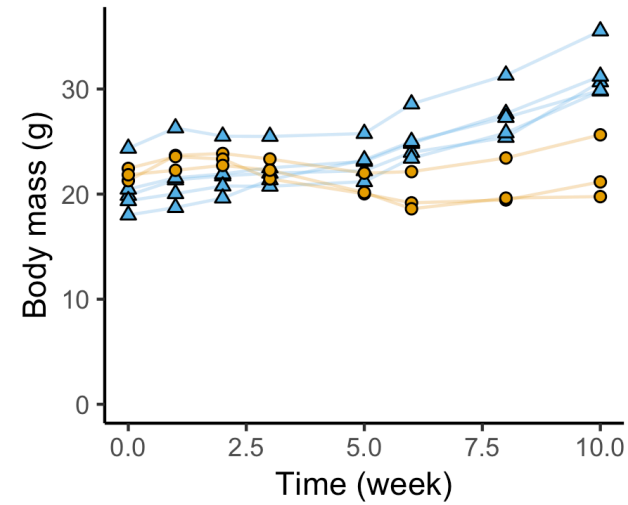


Better approach: build a model

- First: understand your data
- Build a model and reduce time-course curves to just one number
- Do a t-test or similar test on these numbers
- Very simple: area under each curve
- This gives us 4 vs. 3 areas



Compare area under the curve



Chi-square or G-test vs. ANOVA

	WT	KO1	KO2	KO3
G1	50, 54, 48	61, 75, 69	78, 77, 80	43, 34, 49
S	172, 180, 172	175, 168, 166	162, 167, 180	178, 173, 168
G2	55, 50, 63	45, 41, 38	47, 49, 43	59, 50, 45

Fisher's test / Chi-square test / G-test

Experiment outcome: category

Table contains counts

	English	Scottish	Welsh	N. Irish
White	19.1, 20, 21	22.3, 21.2, 25.6	18.1, 19.2, 22.7	15.6, 16.7, 15
Black	21.1, 20, 20.5	21.1, 27.5, 23	22.5, 18.5, 19	19.1, 17.7, 13.5
Grey	20, 21, 17	18.6, 20.1, 19.7	15, 18, 22	12, 18.1, 20.3

ANOVA

Experiment outcome: measurement (could be counts)

Table contains measurements

Chi-square test or ANOVA?

Bacterial antibiotic resistance

- Four strains
- Grown in normal medium and two antibiotic concentrations
- Dilution plating, count colonies

	WT	KO1	KO2	KO3
No antibiotic	77, 51, 92	50, 83, 16	70, 111, 78	121, 147, 110
Conc. 1	83, 51, 40	66, 18, 49	95, 109, 52	75, 116, 109
Conc. 2	11, 7, 31	69, 41, 21	85, 51, 60	95, 128, 116

Outcome is measurement, not category
This is not a contingency table!

Use ANOVA

Hand-outs available at
https://dag.compbio.dundee.ac.uk/training/Statistics_lectures.html